

Integrated Volatility Measuring from Unevenly Sampled Observations.

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November 28, 2003

1 Estimators of integrated volatility

Let p_t be logarithmic asset price

$$dp(t) = \mu(t)dt + \sigma(t)dW(t), \quad (1)$$

where $W(t)$ is a standard Brownian Motion, $\mu(t), \sigma(t)$ are random time dependent functions. The diffusion is observed at $\{t_i\}_{i=0}^N$. In this paper, we compare the estimators of integrated volatility $\int_0^T \sigma^2(t)dt$.

1.1 Quadratic variation of evenly sampled observations through linear interpolation

The transaction data which are unevenly spaced, are not directly used. After creating evenly spaced data $\{p(iT/m)\}_{i=0}^m$ from $\{p(t_i)\}_{i=0}^N$ through linear interpolation, the volatility is measured by the following estimator,

$$\hat{\sigma}^2(m) = \sum_{i=1}^m \left(p\left(\frac{iT}{m}\right) - p\left(\frac{(i-1)T}{m}\right) \right)^2. \quad (2)$$

This estimator is downward biased.

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1.2 Fourier estimator of [MM02]

To avoid the interpolation bias, [MM02] proposed the method without any data manipulation by using Fourier series.

$$\hat{\sigma}_F^2 = 2\pi a_0(\sigma^2) \quad (3)$$

where

$$a_0(\sigma^2) = \lim_{n \rightarrow \infty} \frac{\pi}{n-1} \sum_{s=2}^n \frac{1}{2} (a_k^2(dp) + b_k^2(dp)), \quad (4)$$

$$a_k(dp) = \frac{1}{\pi} \int \cos(kt) dp(t), \quad (5)$$

$$b_k(dp) = \frac{1}{\pi} \int \sin(kt) dp(t), \quad (6)$$

and n is Nyquist frequency $N/2$.

1.3 Quadratic variation of unevenly sampled observations

Another method using unevenly sampled observations $\{p(t_i)\}_{i=0}^N$:

$$\hat{\sigma}^2 = \sum_{i=1}^N (p(t_i) - p(t_{i-1}))^2, \quad (7)$$

has a nice property that if $\sup_{i \geq 1} (t_i - t_{i-1}) \rightarrow 0$,

$$\lim_{N \rightarrow \infty} \hat{\sigma}^2 = \int_0^T \sigma^2(t) dt. \quad (8)$$

See e.g. [ABDL03]. This estimator is simple but as efficient as Fourier estimator.

1.4 Monte Carlo simulations

We generate proxy for continuous observation by discretizing following equations with a time step of one second,

$$\begin{aligned} dp(t) &= \mu(t)dt + \sigma(t)dW(t), \\ d \log \sigma_t &= -k \log \sigma_t dt + \gamma dW_t \end{aligned}$$

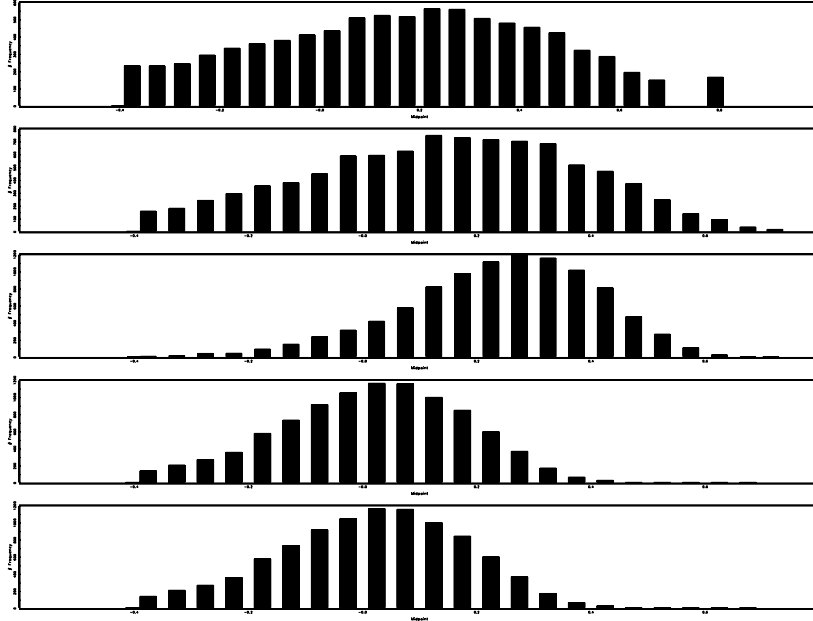


Figure 1: 10min vs 5min vs 2min vs FE vs QV. The distributions are compared with 10,000 replications.

	computational time
Fourier estimator	1116.42''
Quadratic variation	0.25''

Table 1: Computational time (seconds)

where W_s is standard Brownian Motion. The waiting times are drawn from an exponential distribution with mean 45 seconds according to [BR02]. See [ER98] for the modeling of waiting time. Figure 1 reports the distributions of

$$1 - \frac{\hat{\sigma}^2}{\int_0^T \sigma^2(t) dt}.$$

Table report the computational time of FE and QV.

2 Cross-volatility

$$dp_j(t) = \mu_j(t)dt + \sum_{k=1}^d \sigma_{jk}(t)dW_k(t), \quad (9)$$

Volatility matrix is defined by

$$\Omega_{(jk)}(t) = \sum_{i=1}^d \sigma_{ji}\sigma_{ki}.$$

Our target is $\int_0^T \Omega(t) dt$.

2.1 Linear interpolation

$$\hat{\Omega}_{(jk)}(m) = \sum_{i=1}^m \left(p_j \left(\frac{iT}{m} \right) - p_j \left(\frac{(i-1)T}{m} \right) \right) \left(p_k \left(\frac{iT}{m} \right) - p_k \left(\frac{(i-1)T}{m} \right) \right).$$

What occurs on linear interpolation bias?

2.2 Fourier estimator

$$\hat{\Omega}_{F(jk)} = 2\pi a_0(\Omega_{(jk)})$$

where

$$a_0(\Omega_{(jk)}) = \lim_{n \rightarrow \infty} \frac{\pi}{n-1} \sum_{s=2}^n \frac{1}{2} (a_s(dp_j)a_s(dp_k) + b_s(dp_j)b_s(dp_k)), \quad (10)$$

$$a_k(dp_i) = \frac{1}{\pi} \int \cos(kt) dp_i(t), \quad (11)$$

$$b_k(dp_i) = \frac{1}{\pi} \int \sin(kt) dp_i(t), \quad (12)$$

where $n = \lfloor N/2 \rfloor$.

2.3 One-side linear Interpolation

The j th and k th diffusion of (9) are observed at $\{t_i\}_{i=0}^{N_j}$ and $\{t_i\}_{i=0}^{N_k}$ respectively. Define the sequence: $\{t_i\}_{i=0}^{N_{jk}} \equiv \left\{ t : \{t_i\}_{i=0}^{N_j} \cup \{t_i\}_{i=0}^{N_k} \right\}$.

$$\hat{\Omega}_{(jk)}(m) = \sum_{i=1}^{N_{jk}} (p_j(t_i) - p_j(t_{i-1})) (p_k(t_i) - p_k(t_{i-1}))$$

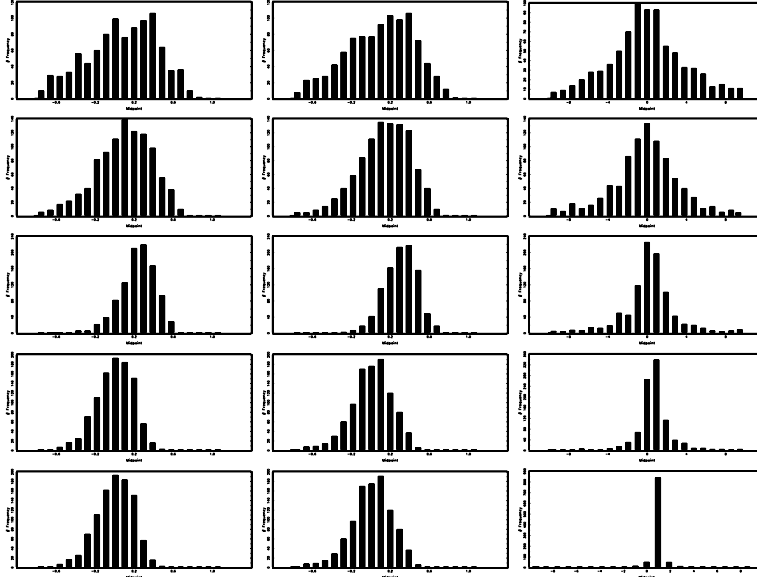


Figure 2: Volatilities and cross-volatility. The distributions are compared with 1,000 replications.

2.4 Monte Carlo simulations

$$\begin{pmatrix} dp_1(t) \\ dp_2(t) \end{pmatrix} = \begin{pmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{pmatrix} \begin{pmatrix} dW_1(t) \\ dW_2(t) \end{pmatrix}$$

$$d\sigma_{jk}(t) = -\kappa_{jk}\sigma_{jk}(t) dt + \gamma_{jk}dW_{jk}(t), \quad j, k = 1, 2.$$

where $\kappa_{jk} = 0.99$ and $\gamma_{jk} = 0.01$ for any j, k .

Figure 2 reports the distributions of

$$1 - \frac{\hat{\Omega}_{(jk)}}{\int_0^T \Omega_{(jk)}(t) dt}.$$

3 Conclusion

Let us use (7) in scalar case. However, we expect that Fourier estimator is good for cross-volatility. There are many remaining works:

- Asymptotic distribution of the estimators.
- Linear interpolation bias correction.
- Long memory.

Acknowledgements

Thank you for your reading. To be continued.

A Fourier estimator of [MM02]

The method will be the following: first compute the Fourier coefficients of dp_i , then obtain a mathematical expression of the Fourier coefficients of Ω_{jk} using the Fourier coefficients of dp_i .

References

- [ABDL03] Torben G. Andersen, Tim Bollerslev, Francis X. Diebold, and Paul Labys, *Modeling and forecasting realized volatility*, *Econometrica* **71** (2003), 579–625.
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- [ER98] Robert F. Engle and Jeffrey R. Russell, *Autoregressive conditional duration: A new model for irregularly spaced transaction data*, *Econometrica* **66** (1998), 1127–1162.
- [MM02] Paul Malliavin and Maria Elvira Mancino, *Fourier series method for measurement of multivariate volatilities*, *Finance and Stochastics* **6** (2002), 49–61.